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Supplementary Materials for

Cross-disciplinary evolution of the genomics revolution

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Appendix S1. Author name disambiguation.

An important challenge in the career dataset is name disambiguation for authors and co-authors. Ambivalence primarily stems from the inconsistent use of suffixes and middle names or initials. Within the collaboration profile of \mathcal{F}_i , there are J_0^i raw name strings, $Name_j^i$, indexed by *j*. The following steps outline the disambiguation procedure we applied to address name conflicts:

A. Clean last names: Remove strings at the end of $Name_j^i$ that are not last names, and which may not be consistently listed for pollinator *j* across the profile of \mathcal{F}_i - e.g., "Jr.", "III", and the like. At the end of this removal process, each pollinator's name string $Name_j^i$ would ideally consist of a first name string FN_i^i , possibly a middle name string MN_j^i , and a last name string LN_j^i .

B. Disambiguate middle initial strings within each \mathcal{F} **profile:** Within the profile of each \mathcal{F}_i , search for inconsistencies in the use of $MN_{j?}^i$. For example, sometimes pollinator *j*? may be listed as *Amanda M Price*, some other times as *Amanda Price*, and yet some other times as *Amanda Miranda Price*. In this example, the last name string $LN_{j?} = Price$ and the first name string $FN_{j?} = Amanda$ are consistent. However, the middle name string set {_, M, Miranda} introduces ambiguity, as it includes instances of no middle name $MN_{j?}^i \equiv \emptyset$, middle initial $MN_{j?}^i \equiv J$, and full middle name $MN_{j?}^i \equiv Jname$, compatible with the middle initial instances. Depending on the type of middle name ambiguity for pollinator *j*?, apply the following rules:

- If the middle name ambiguity set for pollinator *j*? has instances of no middle name and middle initial only $\{\emptyset, \mathcal{I}\}$, then transform all name instances of pollinator *j*? to $\langle FN_j^i \ \mathcal{I} \ LN_j^i \rangle$, firmly assigning them the *j* index.
- If the middle name ambiguity set for pollinator *j*? has instances of no middle name, middle initial, and full middle name, compatible with the middle initial instances $\{\emptyset, J, Jname\}$, then transform all name instances of pollinator *j*? to $\langle FN_i^i \ Jname \ LN_j^i \rangle$, firmly assigning them the *j* index.
- If the middle name ambiguity set for pollinator j? has instances of no middle name and two different middle initials {Ø, J1, J2}, then check if (FN_{j?} LN_{j?}) and (FN_{j?} J1 LN_{j?}) are co-authors in the same paper within the F profile i. If they are, then transform (FN_{j?} LN_{j?}) to (FN_{j2} J2 LN_{j2}), assigning to the consolidated subset a j₂ index. If the first test fails, then check if (FN_{j?} LN_{j?}) and (FN_{j?} J2 LN_{j?}) are co-authors in the same paper within the F profile i. If they are, then transform (FN_{j?} J2 LN_{j?}) are co-authors in the same paper within the F profile i. If they are, then transform (FN_{j?} J2 LN_{j?}) are co-authors in the same paper within the F profile i. If they are, then transform (FN_{j?} LN_{j?}) to (FN_{j1} J1 LN_{j1}), assigning to the consolidated subset a j₁ index. If both tests fail, then compare the co-authors among (FN_{j?} LN_{j?}), (FN_{j?} J1 LN_{j?}), and (FN_j J2 LN_{j?}), and

 $\langle FN_{j?}^i \ \mathcal{J}2 \ LN_{j?}^i \rangle$ within the \mathcal{F} profile *i*. Transform the no middle name instances to the middle name variety with which it shares more co-authors.

C. Disambiguate pollinators across \mathcal{F} **profiles:** Let *j* and *j'* be pollinators in \mathcal{F} profiles *i* and *i'*, respectively. Check if *j* and *j'* are likely the same person, $j \equiv j'$, in order to establish (or not) a mediated association link between *i* and *i'*. Depending on the type of ambiguity, apply the following rules:

• If the ambiguous pollinators have the same first names, $FN_j^i = FN_{j\prime}^{i\prime}$, and the same last names, $LN_j^i = LN_{j\prime}^{i\prime}$, and the middle name ambiguity set consists of no middle name plus at least two different middle names { \emptyset , Jname1, Jname2, ... }, then compare the common co-authors among

the no middle name instances and the other instances, assigning the no middle name instance to the case with which it shares the most common co-authors.

- If the first name of a pollinator *j* is hyphenated $FN_j^i \equiv FN1_j^i FN2_j^i$, and has only 2 letters i.e., *FN*1 and *FN*2 are one letter, check for any other pollinator *j'* that has hyphenated first name with the same first letter $FN1_j^i$ and the first letter after the hyphen starts with $FN2_j^i$. Then transform the *j* pollinator to *j'* who has the longest such hyphenated first name.
- If the first name of a pollinator *j* is hyphenated $FN_j^i \equiv FN1_j^i FN2_j^i$, check for any other pollinator *j'* that has $FN_{j'}^{i'} = FN1_j^i$, $\mathcal{I}' = FN2_j^i$, and $LN_{j'}^{i'} = LN_j^i$. If such a pollinator *j'* does exist and shares at least one common co-author with *j*, then transform the *j* pollinator to *j'*, assigning the second part of her/his original hyphenated first name to be her/his middle name.
- If the name of a pollinator *j* has only two letters $FN_j^i \equiv \mathcal{L}1_j^i \mathcal{L}2_j^i$, check for any other pollinator *j'* that has $FN_{j'}^{i'} = \mathcal{L}1name_j^i$, $\mathcal{I}' = \mathcal{L}2name_j^i$, and $LN_{j'}^{i'} = LN_j^i$. If such a pollinator *j'* does exist and shares at least one common co-author with *j*, then transform the *j* pollinator to *j'*.

Appendix S2. Connectivity of the \mathcal{F} network.

How does the \mathcal{F} network depend on the direct \mathcal{F} connectivity? To investigate this, we randomly removed a fraction q of the links, incrementing q over the range [0,1], and monitoring the effect on the network's giant and non-giant components. This method of random link removal is drawn from the theory of phase transitions in the connectivity of networks (62,63). For each q, we performed the link percolation 40 times and reported the mean and standard deviation of the following network connectivity descriptors:

Giant component size:

For the \mathcal{F} collaboration network, the initial size of the largest connected component (aka giant component) is $S_G(q = 0) = 3,869$, meaning that 321 \mathcal{F} nodes are initially disconnected from the giant component. Figure S1a shows the ratio $S_G(q)/S_G(q=0)$ as a function of q, demonstrating the robustness of the collaboration network - even after 80% of the links are removed, roughly 60% of the \mathcal{F} are still connected within the network. Of course, the fragmentation of the network depends on how we remove the links. We compared the results for random uniform removal of links and for random removal according to the weight W_{ii} . For each W_{ii} , definition, we removed the links according to increasing weight and also according to the inverted weight $W_{ii\prime}^r = \max_{ii\prime} [W_{ii\prime}] - W_{ii\prime}$ ('reverse'). We used three definitions for the link weights W_{ii} : (a) $W_{ii} \equiv \max[PR_i, PR_i]$, where PR_i and PR_i are the PageRank centralities of node i and i', respectively, using the common damping factor 0.85; (b) $W_{ii} \equiv \max[B_i, B_{i'}]$, where B_i and $B_{i'}$ are the betweenness centralities of nodes i and i', representing the number of shortest paths in the network that traverse i and i', respectively; (c) $W_{ii'} \equiv O_{ii'}$, where $O_{iii} \in [0,1]$ is the overlap fraction in the first-degree neighbors of nodes *i* and *i'*, calculated as $O_{iii} = O_{iii}$ $s_{ii'}/[(k_i - 1) + (k_{i'} - 1) - s_{ii'}]$, where $s_{ii'}$ is the number of shared first-degree neighbors, and k_i and k_{ii} are the degrees of nodes i and i', respectively (64). Consistent with expectations, the link removal methods that exhibited the sharpest fragmentation were PR_{ii} , and B_{ii} .

Susceptibility to fragmentation:

For each q we calculated the size S_i of all the N_q fragments, where by definition $S_G(q) = \max_i (S_i(q))$. The severity of the fragmentation (percolation) process can be further illustrated by analyzing the fragment size distribution $P(S_i)$, i.e., by calculating the distribution's second moment $\sigma_q^2 = \sum_{i|S_i < S_G}^{N_q-1} S_i^2 P(S_i)$. By construction, σ_q^2 does not include the giant component S_G . The fluctuation scale σ_q^2 diverges when the network shatters into pieces of varying sizes. Indeed, fig. S1b shows how the network's susceptibility to fragmentation peaks - depending on the link removal weights – when there is a precipitous drop in the connectivity of the giant component (fig. S1a). The fragmentation peak is associated with the critical point of the network, and is achieved at a smaller q value when the links associated with the most central \mathcal{F} are removed first (blue and black curves in fig. S1b).



Fig. S1. Robustness of the \mathcal{F} network with respect to link removal. (a) The ratio $S_G(q)/S_G(q = 0)$ measures the size of the largest remaining fragment $S_G(q)$, relative to the size of the initial giant component $S_G(q = 0)$. The slow decay until q = 0.6 indicates that this network is robust to variation in the connectivity of scholars. For a given q, we repeated the fragmentation process 40 times, and plotted the error bars to indicate the mean and standard deviation. (b) Detection of the critical point at which the college disassociates. For each q we also monitor the size S_i of all the N_q disconnected network fragments, where by definition $S_G(q) = \max_i(S_i(q))$. As a limiting example, complete disassociation occurs for q = 1 (all links removed), corresponding to a completely disconnected ensemble of nodes with $N_q = 4,190$ and $S_i = 1$ for all *i*. The fluctuation scale of the fragmentation process is illustrated by the variation of the fragment size distribution, σ_q^2 , which diverges when the network 'shatters' into pieces of highly variable sizes. The peak in σ_q^2 signals the onset of the shattering process.



Fig. S2. \mathcal{F} network distributions for direct and mediated associations. for a biology and b. computing. Each panel shows the frequency distribution (counts) of faculty \mathcal{F} with a given link degree counting the number of links for a given node, $d_i \equiv \mathcal{C}_i^D(t)$, within a particular definition of the \mathcal{F} network (vertical lines indicate distribution means). The direct subnetworks only include direct links, which are established whenever two \mathcal{F} collaborate on at least one publication. The mediated subnetworks only include indirect links between two \mathcal{F} who have both collaborated with a common pollinator (i.e., are associated via triadic closure - see Fig. 1). On average, 97% in biology and 92% in computing are *pollinator* co-authors, i.e., researchers not included in the \mathcal{F} set. The significantly different scale of the degree distributions demonstrates the connectivity power of the pollinators within the invisible college.



Fig. S3. Three perspectives on the centrality of \mathcal{F}_i in the direct collaboration network. Shown is the giant connected component of the faculty network \mathcal{F} using all data up to 2015. The nodes and links across each network are fixed, only the node sizes vary according to the indicated centrality measure: (a) degree \mathcal{C}_i^D , (b) PageRank \mathcal{C}_i^{PR} , (c) betweenness \mathcal{C}_i^B . Notably, the most central \mathcal{F}_i according to each of the three measures is Eric Lander, one of the leaders of the HGP.



Fig. S4. Evolution of the nongiant components in the \mathcal{F} network. Green and magenta nodes represent faculty \mathcal{F}_i with $BIO_{\mathcal{F}}$ and $CS_{\mathcal{F}}$ affiliation, respectively; black nodes represent faculty \mathcal{F}_i that by time *t* collaborated with at least one faculty from the opposite department and thus joined the $XD_{\mathcal{F}}$ group.



Fig. S5. Distribution of normalized citation impact by departmental affiliation and time period. Probability distribution P(z|s, t) calculated by separating the publications of the career dataset into subsets according to the departmental affiliation of \mathcal{F}_i and publication year. Shown are the empirical distribution (red bins) and baseline normal distribution N(0,1) (blue curve), which demonstrates the time-independence of the normalized citation impact variable.

Table S1. Set of 155 biology and computing departments in the United States.for the career dataset. Ranks are per the 2014 U.S. News & World Report.

| Rank | Biology Departments | Rank | Computing Departments |
|----------|---|----------|--|
| 1 | Harvard University Massachusetts Institute of Technology | 1 | Carnegie Mellon University Massachusetts Institute of Technology |
| 1 | Stanford University | 1 | Stanford University |
| 4 | University of California, Berkeley | 1 | University of California, Berkeley |
| 5 | California Institute of Technology | 5 | University of Illinois, Urbana Champaign |
| 5 | Johns Hopkins University University of California, San Francisco | 6 | Cornell University University of Washington |
| 7 | Yale University | 8 | Princeton University |
| 9 | Princeton University | 9 | Georgia Institute of Technology |
| 9 | Scripps Research Institute | 9 | University of Texas, Austin |
| 11 | Cornell University | 11 | California Institute of Technology |
| 11 | Washington University in St. Louis | 11 | University of California I os Angeles |
| 14 | Columbia University | 13 | University of Michigan, Ann Arbor |
| 14 | Rockefeller University | 15 | Columbia University |
| 14 | University of California, San Diego | 15 | University of California, San Diego |
| 14 | University of Wisconsin Madison | 15 | University of Maryland, College Park Harvard University |
| 19 | University of California, Davis | 19 | University of Pennsylvania |
| 19 | University of California, Los Angeles | 20 | Brown University |
| 19 | University of Michigan, Ann Arbor | 20 | Purdue University, West Lafayette |
| 19 | University of Pennsylvania | 20 | Rice University |
| 19 | University of Texas Southwestern Medical Center | 20 | Vale University |
| 25 | Baylor College of Medicine | 20 | Duke University |
| 26 | Cornell University (Weill) | 25 | University of Massachusetts, Amherst |
| 26 | Northwestern University | 25 | University of North Carolina, Chapel Hill |
| 26 | University of North Carolina, Chapel Hill | 28 | Johns Hopkins University |
| 26 | Emory University | 29 | New YORK University Pennsylvania State University University Park |
| 30 | University of Colorado, Boulder | 29 | University of California, Irvine |
| 30 | University of Illinois, Urbana Champaign | 29 | University of Minnesota, Twin Cities |
| 30 | University of Texas, Austin | 29 | University of Virginia |
| 34 | Brown University | 34 | Northwestern University |
| 34 34 | University of California Irvine | 54 34 | Rutgers The State University of New Jersey |
| 34 | University of Minnesota, Twin Cities | 34 | University of California, Davis |
| 38 | Case Western Reserve University | 34 | University of California, Santa Barbara |
| 38 | Dartmouth College | 34 | University of Chicago |
| 38 | Mayo Medical School | 40 | Dartmouth College |
| 30 42 | Carnegie Mellon University | 40 | Texas A&M University College Station |
| 42 | Icahn School of Medicine at Mount Sinai | 40 | University of Arizona |
| 42 | Ohio State University | 40 | University of Colorado, Boulder |
| 42 | Pennsylvania State University, University Park | 40 | University of Utah |
| 42 | Kice University | 40 | Virginia Iech Washington University in St. Louis |
| 42 | University of Georgia | 40 | Arizona State University |
| 42 | University of Pittsburgh | 48 | Boston University |
| 50 | Michigan State University | 48 | North Carolina State University |
| 50 | University of California, Santa Barbara | 48 | University of Florida |
| 50 50 | University of Massachusetts Medical Center | 52 52 | Indiana University, Bloomington Peneselaer Polytechnic Institute |
| 50 | Yeshiya University (Einstein) | 52 | University of Pittsburgh |
| 55 | Arizona State University | 52 | University of Rochester |
| 55 | Brandeis University | 56 | Michigan State University |
| 55 | Georgia Institute of Technology | 56 | University of California, Riverside |
| 55 | Stony Brook University SUNY | 56 | Vanderbilt University |
| 55 | University of California, Santa Cruz | 60 | Northeastern University |
| 55 | University of Florida | 60 | University of Illinois, Chicago |
| 55 | University of Iowa | 60 | University of Notre Dame |
| 55 | University of Maryland, College Park | 63 | Iowa State University |
| 55 | University of Oregon | 63 | University at Bullalo, SUN I |
| 55 | University of Southern California | 63 | University of Oregon |
| 55 | University of Utah | 67 | George Mason University |
| 68 | New York University | 67 | Oregon State University |
| 68 68 | Oregon Health and Science University | 67 70 | Syracuse University |
| 68 | Tufts University | 70 | College of William and Mary |
| 68 | University of California. Riverside | 70 | Colorado State University |
| 68 | University of Kansas | 70 | Naval Postgraduate School |
| 68 | University of Rochester | 70 | New York University |
| | | 70 | Tutts University |
| | | 70 | University of Delaware |
| | | 70 | University of Maryland, Baltimore County |
| | | 70 | University of Nebraska, Lincoln |
| | | 70 | University of Tennessee, Knoxville |
| | | 70 | University of Texas, Dallas |
| | | 70 | Washington State University |

Table S2. Career data set: Pooled cross-sectional model. The dependent variable is career achievement, measured as the natural logarithm of the Google Scholar citations, $\ln C_i$ as of 2017. The regression model is specified in Eq. (1) and estimated using standard OLS; there are 4,190 \mathcal{F}_i (observations) for the pure CV model and 3,900 observations for the other two models that include network attributes, as in these cases we exclude from consideration disconnected \mathcal{F}_i nodes. Natural logs were used to obtain variables that are approximately normally distributed. Thus, when the independent variable enters in ln, then β corresponds to the % change in C_i following a 1% change in the independent variable; in the case of the cross-disciplinarity fraction, β_{χ} represents the % change in C_i following a 0.01 shift increase in χ_i . The first column cluster shows the estimates using only standard CV variables. The combined CV + Network model demonstrates that \mathcal{F}_i with larger χ_i correlate with higher net citation impact. For the combined model we also report the standardized beta coefficients – useful for comparing the relative strength of covariates within the regression. Standard errors were calculated using the clustered sandwich estimator, clustering on \mathcal{F}_i age-cohort $y_{i,5}^0$ (based on 14 non-overlapping 5-year career birth year groups, e.g., 1940-1944, 1945-1950, etc.) to account for within-age-cohort correlation. Y indicates additional fixed effects included in the regression model.

| | CV | | CV + Network | | CV + Network [Standardized] | |
|---|----------------|---------|----------------|---------|-----------------------------|---------|
| CV parameters | | | | | | |
| Departmental rank, β_r | -0.052^{***} | (0.006) | -0.047^{***} | (0.005) | -0.056^{***} | (0.006) |
| Productivity (<i>h</i> -index), β_h | 1.857^{***} | (0.020) | 1.866^{***} | (0.022) | 1.236^{***} | (0.015) |
| Total NSF funding, $\beta_{\$1}$ | -0.005 | (0.003) | -0.005 | (0.003) | -0.036 | (0.020) |
| # of NSF grants, β_{N1} | 0.024 | (0.013) | 0.013 | (0.014) | 0.015 | (0.015) |
| Total NIH funding, $\beta_{\$2}$ | 0.016^{***} | (0.003) | 0.014^{***} | (0.002) | 0.082^{***} | (0.014) |
| # of NIH grants, β_{N2} | -0.067^{***} | (0.015) | -0.061^{***} | (0.012) | -0.068^{***} | (0.014) |
| Network parameters | | | | | | |
| PageRank Centrality, $\beta_{\mathscr{C}^{PR}}$ | | | 0.041 | (0.019) | 0.026 | (0.012) |
| Cross-disciplinarity, β_{χ} | | | 0.571^{***} | (0.073) | 0.085^{***} | (0.011) |
| Discipline (\mathcal{O}) dummy | Y | | Y | | Y | |
| 5-year cohort ($y_{i,5}^0$) dummy | Y | | Y | | Y | |
| Constant | 1.492^{***} | (0.087) | 1.668^{***} | (0.226) | 7.609^{***} | (0.009) |
| | 4,190 | | 3,900 | | 3,900 | |
| adj. R^2 | 0.883 | | 0.882 | | 0.882 | |

Standard errors in parentheses

* $p \le .05$, ** $p \le .01$, *** $p \le .001$

Table S3. Career data set: Pooled cross-sectional model—robustness check. Parameter estimates for variants of the 'CV + Network' pooled cross-sectional models reported in table S2: (a) Model with PageRank centrality. (b) Model with betweenness centrality. (c) Model with degree centrality; (d) Model without the number of grants variables; (e) Model without the departmental rank variable. Results are not significantly different with respect to the primary covariate of interest, that is, cross-disciplinarity (β_{χ}).

| | (a) | (b) | (c) | (d) | (e) |
|---|--------------------|-----------------|-----------------|-----------|----------------|
| | \mathscr{C}^{PR} | \mathscr{C}^B | \mathscr{C}^D | BNI, BN2 | β_r |
| CV parameters | | | | | |
| Departmental rank, β_r | -0.047^{***} | -0.042*** | -0.044^{***} | -0.046*** | |
| | (0.005) | (0.005) | (0.005) | (0.005) | |
| Productivity (<i>h</i> -index), β_h | 1.866^{***} | 1.901*** | 1.848*** | 1.862*** | 1.892^{***} |
| | (0.022) | (0.024) | (0.024) | (0.020) | (0.022) |
| Total NSF funding, $\beta_{\$1}$ | -0.005 | -0.004 | -0.005 | -0.003 | -0.005 |
| | (0.003) | (0.004) | (0.003) | (0.002) | (0.003) |
| # of NSF grants, β_{N1} | 0.013 | 0.009 | 0.007 | | 0.006 |
| | (0.014) | (0.020) | (0.014) | | (0.015) |
| Total NIH funding, $\beta_{\$2}$ | 0.014^{***} | 0.014*** | 0.014^{***} | 0.003* | 0.013^{***} |
| | (0.002) | (0.003) | (0.002) | (0.001) | (0.003) |
| # of NIH grants, β_{N2} | -0.061^{***} | -0.065*** | -0.062^{***} | | -0.059^{***} |
| | (0.012) | (0.013) | (0.012) | | (0.014) |
| Network parameters | | | | | |
| PageRank centrality, $\beta_{\mathscr{C}^{PR}}$ | 0.041 | | | 0.042 | 0.057^{*} |
| | (0.019) | | | (0.019) | (0.020) |
| Betweeness centrality, $\beta_{\mathscr{C}^B}$ | | -0.000 | | | |
| | | (0.006) | | | |
| Degree centrality, $\beta_{\mathscr{C}^D}$ | | | 0.052^{**} | | |
| | | | (0.016) | | |
| Cross-disciplinarity, eta_{χ} | 0.571^{***} | 0.562*** | 0.530^{***} | 0.579*** | 0.555^{***} |
| | (0.073) | (0.054) | (0.072) | (0.073) | (0.076) |
| Discipline (\mathcal{O}) dummy | Y | Y | Y | Y | Y |
| 5-year cohort ($y_{i,5}^0$) dummy | Y | Y | Y | Y | Y |
| Constant | 1.668^{***} | 1.192*** | 1.293^{***} | 1.671*** | 1.579^{***} |
| | (0.226) | (0.083) | (0.069) | (0.226) | (0.225) |
| n | 3900 | 3387 | 3900 | 3900 | 3900 |
| adj. R^2 | 0.882 | 0.873 | 0.883 | 0.882 | 0.881 |

Standard errors in parentheses, listed below coefficient estimate.

* $p \le .05$, ** $p \le .01$, *** $p \le .001$

Table S4. Career data set: Panel model on all faculty \mathcal{F} . Each column cluster reports the estimated coefficients for a specific model in which the dependent variable is the normalized citation impact of an individual article, $z_{i,p}$ belonging to faculty \mathcal{F}_i - see Eq. (4). The first two column clusters correspond to a panel regression without \mathcal{F}_i fixed effects, whereas the last two column clusters correspond to a panel regression with \mathcal{F}_i fixed effects. Estimates in the second and fourth column clusters are calculated using standardized variables, where each 'beta' coefficient indicates the change in $z_{i,p}$ associated with a one standard deviation shift in the corresponding independent variable. The model without fixed effects incorporates time-independent author-level characteristics, i.e., adding to the specification of Eq. (2) the additional terms $[\beta_{\mathcal{C}_i}^{PR} \ln \mathcal{C}_i^{PR} + \beta_{\lambda} \ln \lambda_i + D(\mathcal{F}_i)]$. This is the reason why we only analyzed the 3,900 scholars connected within the network for which \mathcal{C}_i^{PR} is defined; note that these additional variables are absorbed into β_i in the fixed effects model. The additional connectivity variable λ_i is the fraction of the total pollinators that are 'bridge' pollinators. Robust standard errors are shown in parenthesis, and X denotes time-independent variables absorbed by the fixed effects model. Y indicates additional fixed effects included in the regression model.

| | No Fixed Effects | No Fixed Effects | Fixed Effects | Fixed Effects |
|---|------------------|------------------|---------------|----------------|
| | | [Standardized] | | [Standardized] |
| Publication characteristics | | | | |
| # of co-authors, β_a | 0.284*** | 0.189*** | 0.312*** | 0.208*** |
| | (0.00718) | (0.00479) | (0.00547) | (0.00365) |
| Career age, β_{τ} | -0.00547*** | -0.0564*** | -0.00949*** | -0.0978*** |
| | (0.000919) | (0.00947) | (0.00182) | (0.0187) |
| Cross-disciplinary indicator, β_I | 0.126*** | 0.126*** | 0.145*** | 0.145*** |
| | (0.0341) | (0.0341) | (0.0235) | (0.0235) |
| Network characteristics | | | | |
| PageRank centrality, $\beta_{\mathscr{C}^{PR}}$ | 0.0440** | 0.0284** | Х | Х |
| | (0.0142) | (0.00920) | | |
| Bridge fraction, β_{λ} | 0.334*** | 0.129*** | Х | Х |
| | (0.0256) | (0.0099) | | |
| Discipline (\mathcal{F}) dummy | 0.00790 | 0.00790 | Х | Х |
| | (0.0139) | (0.0139) | | |
| Constant | 0.451** | 0.151 | -0.293*** | -0.0653** |
| | (0.142) | (0.102) | (0.0528) | (0.0202) |
| Year dummy | Y | Y | Y | Y |
| n | 413,565 | 413,565 | 413,565 | 413,565 |
| adj. R^2 | 0.055 | 0.055 | 0.036 | 0.036 |

Standard errors in parentheses below estimate.

* $p \le 0.05$, ** $p \le 0.01$, *** $p \le 0.001$

Table S5. Career data set: Panel model on the $XD_{\mathcal{F}}$ **faculty.** Robustness check of panel model without and with fixed effects, implemented using only the 1,247 \mathcal{F}_i with orientation $\mathcal{O}(\mathcal{F}_i) = XD_{\mathcal{F}}$.

| | No Fixed Effects | No Fixed Effects [Standardized] | Fixed Effects | Fixed Effects [Standardized] |
|--|------------------|------------------------------------|---------------|---------------------------------|
| Publication characteristics | | | | |
| # of co-authors, β_a | 0.329*** | 0.220*** | 0.351*** | 0.234*** |
| | (0.0123) | (0.00821) | (0.00880) | (0.00588) |
| Career age, β_{τ} | -0.00499** | -0.0514** | -0.00616* | -0.0635* |
| | (0.00181) | (0.0187) | (0.00253) | (0.0261) |
| Cross-disciplinary indicator, β_I | 0.109*** | 0.109*** | 0.112*** | 0.112*** |
| | (0.0328) | (0.0328) | (0.0234) | (0.0234) |
| Network characteristics | | | | |
| Author centrality, $\beta_{\mathscr{C}}$ | 0.0526* | 0.0340* | Х | Х |
| | (0.0265) | (0.0171) | | |
| Bridge fraction, β_{λ} | 0.319*** | 0.124*** | Х | Х |
| | (0.0493) | (0.0191) | | |
| Discipline (\mathcal{F}) dummy | -0.0383 | -0.0383 | Х | Х |
| | (0.0256) | (0.0256) | | |
| Constant | 0.217 | -0.0773 | -0.409*** | -0.0685* |
| | (0.236) | (0.157) | (0.0778) | (0.0324) |
| Year dummy | Y | Y | Y | Y |
| n | 166,621 | 166,621 | 166,621 | 166,621 |
| adj. R^2 | 0.067 | 0.067 | 0.049 | 0.049 |

Standard errors in parentheses below estimate.

* $p \le 0.05$, ** $p \le 0.01$, *** $p \le 0.001$

Table S6. Career data set: Panel model on the $XD_{\mathcal{F}}$ **faculty with matched pairs.** Robustness check of panel model without and with fixed effects, implemented using only the 53 \mathcal{F}_i with orientation $\mathcal{O}(\mathcal{F}_i) = XD_{\mathcal{F}}$ who have at least 10 matched pairs of publications. Where possible, we matched each p with $I_{i,p}^{XD} = 1$ with a publication with $I_{i,p}^{XD} = 0$ from the same \mathcal{F}_i , having published within two years from each other, and featuring number of co-authors a_p that do not differ more than 20%.

| | No Fixed Effects | No Fixed Effects [Standardized] | Fixed Effects | Fixed Effects [Standardized] |
|--|------------------|------------------------------------|---------------|---------------------------------|
| Publication characteristics | | | | |
| # of co-authors, β_a | 0.511*** | 0.375*** | 0.508*** | 0.373*** |
| | (0.0451) | (0.0331) | (0.0453) | (0.0333) |
| Career age, β_{τ} | -0.00221 | -0.0229 | 0.0474*** | 0.491*** |
| | (0.00582) | (0.0603) | (0.00502) | (0.0519) |
| Cross-disciplinary indicator, β_I | 0.133** | 0.133** | 0.135** | 0.135** |
| | (0.0469) | (0.0469) | (0.0471) | (0.0471) |
| Network characteristics | | | | |
| Author centrality, $\beta_{\mathscr{C}}$ | 0.217* | 0.143* | Х | Х |
| | (0.0980) | (0.0648) | | |
| Bridge fraction, β_{λ} | 0.748** | 0.222** | Х | Х |
| | (0.229) | (0.0682) | | |
| Discipline (\mathcal{F}) dummy | -0.208 | -0.208 | Х | Х |
| | (0.116) | (0.116) | | |
| Constant | -0.414 | -1.538*** | -2.268*** | -0.307*** |
| | (0.748) | (0.171) | (0.0980) | (0.0862) |
| Year dummy | Y | Y | Y | Y |
| n | 1930 | 1930 | 1930 | 1930 |
| adj. R^2 | 0.252 | 0.252 | 0.093 | 0.093 |

Standard errors in parentheses below estimate.

* $p \leq 0.05,$ ** $p \leq 0.01,$ **
* $p \leq 0.001$